

# Multiple Linear Regression

CIVL 7012/8012



# Multiple Regression Analysis (MLR)

- Allows us to explicitly control for many factors those simultaneously affect the dependent variable
- This is important for
  - examining theories
  - assessing various policies of independent variables
- MLR can accommodate many independent variables that may be correlated with the dependent variable we can infer causality.
  - In such instances simple regression analysis may be misleading or underestimate the model strength

# MLR Motivation

- Incorporate more explanatory factors into the model
- Explicitly hold fixed other factors that otherwise would be in  $u$
- Allow for more flexible functional forms
- Can take as many as independent variables

# MLR Notation

- Explains “y” in terms of  $x_1, x_2, \dots, x_k$



Intercept

Slope parameters

Dependent variable,  
explained variable,  
response variable,...

Independent variables,  
explanatory variables,  
regressors,...

Error term,  
disturbance,  
unobservables,...

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + u$$

# MLR Example-1

- Wage equation

Now measures effect of education explicitly holding experience fixed

$$wage = \beta_0 + \beta_1 educ + \beta_2 exper + u$$

← All other factors...

Hourly wage

Years of education

Labor market experience



# MLR Example-2

- Average test score, student spending, and income

$$avgscore = \beta_0 + \beta_1 expend + \beta_2 avginc + u$$

Average standardized  
test score of school

Per student spending  
at this school

Average family income  
of students at this school

Other factors

# MLR Example-3

- Example: Family income and family consumption

$$\text{Family consumption} = \beta_0 + \beta_1 \text{Family income} + \beta_2 \text{Family income squared} + \text{Other factors}$$

The equation is annotated with red arrows pointing from text labels to terms in the equation:  $cons = \beta_0 + \beta_1 inc + \beta_2 inc^2 + u$ . The label "Family consumption" points to  $cons$ , "Family income" points to  $\beta_1 inc$ , "Family income squared" points to  $\beta_2 inc^2$ , and "Other factors" points to  $u$ . The term  $u$  is circled in red.

- Model has two explanatory variables: income and income squared
- Consumption is explained as a quadratic function of income
- One has to be very careful when interpreting the coefficients:

By how much does consumption increase if income is increased by one unit?

$$\frac{\partial cons}{\partial inc} = \beta_1 + 2\beta_2 inc$$

Depends on how much income is already there

The partial derivative equation is annotated with red arrows: one from the text "By how much does consumption increase if income is increased by one unit?" to the derivative symbol, and another from the text "Depends on how much income is already there" to the  $inc$  term in the equation.

# MLR Example-4

- Example: CEO salary, sales and CEO tenure

$$\log(\text{salary}) = \beta_0 + \beta_1 \log(\text{sales}) + \beta_2 \text{ceoten} + \beta_3 \text{ceoten}^2 + u$$

Log of CEO salary

Log sales

Quadratic function of CEO tenure with firm

- Model assumes a constant elasticity relationship between CEO salary and the sales of his or her firm
- Model assumes a quadratic relationship between CEO salary and his or her tenure with the firm
- Meaning of linear regression
  - The model has to be linear in the parameters (not in the variables)



# Parallels with Simple Regression

- $\beta_0$  is still the intercept
- $\beta_1$  to  $\beta_k$  all called slope parameters
- $u$  is still the error term (or disturbance)
- Still need to make a zero conditional mean assumption, so now assume that
- $E(u | x_1, x_2, \dots, x_k) = 0$
- Still minimizing the sum of squared residuals, so have  $k+1$  first order conditions

# Interpreting Multiple Regression (1)

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \dots + \hat{\beta}_k x_k, \text{ so}$$

$$\Delta \hat{y} = \Delta \hat{\beta}_1 x_1 + \Delta \hat{\beta}_2 x_2 + \dots + \Delta \hat{\beta}_k x_k,$$

so holding  $x_2, \dots, x_k$  fixed implies that

$$\Delta \hat{y} = \Delta \hat{\beta}_1 x_1, \text{ that is each } \beta \text{ has}$$

*a ceteris paribus* interpretation

# Interpreting Multiple Regression (2)

- Interpretation of the multiple regression model

$$\beta_j = \frac{\partial y}{\partial x_j}$$

By how much does the dependent variable change if the j-th independent variable is increased by one unit, holding all other independent variables and the error term constant

- The multiple linear regression model manages to hold the values of other explanatory variables fixed even if, in reality, they are correlated with the explanatory variable under consideration
- Ceteris paribus-interpretation
- It has still to be assumed that unobserved factors do not change if the explanatory variables are changed

# MLR Estimation (1)

## OLS Estimation of the multiple regression model

- Random sample

$$\{(x_{i1}, x_{i2}, \dots, x_{ik}, y_i) : i = 1, \dots, n\}$$

- Regression residuals

$$\hat{u}_i = y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \hat{\beta}_2 x_{i2} - \dots - \hat{\beta}_k x_{ik}$$

- Minimize sum of squared residuals

$$\min \sum_{i=1}^n \hat{u}_i^2 \rightarrow \hat{\beta}_0, \hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k$$

Minimization will be carried out by computer



# MLR Estimation (2)

- Estimates can be derived from the first order **conditions** of OLS on any sample of data

- Fitted values and residuals

$$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_{i1} + \hat{\beta}_2 x_{i2} + \dots + \hat{\beta}_k x_{ik}$$

Fitted or predicted values

$$\hat{u}_i = y_i - \hat{y}_i$$

Residuals

- Algebraic properties of OLS regression

$$\sum_{i=1}^n \hat{u}_i = 0$$

Deviations from regression line sum up to zero

$$\sum_{i=1}^n x_{ij} \hat{u}_i = 0$$

Correlations between deviations and regressors are

$$\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}_1 + \dots + \hat{\beta}_k \bar{x}_k$$

Sample averages of y and of the regressors lie on regression

# Goodness-of-fit (1)

We can think of each observation  $n$  as being made up of an explained part, and an unexplained part,

$y_i = \hat{y}_i + \hat{u}_i$  We then define the following :

$\sum (y_i - \bar{y})^2$  is the total sum of squares (SST)

$\sum (\hat{y}_i - \bar{y})^2$  is the explained sum of squares (SSE)

$\sum \hat{u}_i^2$  is the residual sum of squares (SSR)

Then  $SST = SSE + SSR$

# Goodness-of-fit (2)

- ◆ How do we think about how well our sample regression line fits our sample data?
- ◆ Can compute the fraction of the total sum of squares (SST) that is explained by the model, call this the R-squared of regression
- ◆  $R^2 = SSE/SST = 1 - SSR/SST$



# More about $R$ -squared

- $R^2$  can never decrease when another independent variable is added to a regression, and usually will increase
- Because  $R^2$  will usually increase with the number of independent variables, it is not a good way to compare models





# Assumptions on MLR (1)



- Standard assumptions for the multiple regression model
- Assumption MLR.1 (Linear in parameters)

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + u$$

In the population, the relationship between  $y$  and the explanatory variables is linear

- Assumption MLR.2 (Random sampling)

$$\{(x_{i1}, x_{i2}, \dots, x_{ik}, y_i) : i = 1, \dots, n\}$$

The data is a random sample drawn from the population

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik} + u_i$$

Each data point therefore follows the population equation



# Assumptions on MLR (2)

- **Standard assumptions for the multiple regression model (cont.)**

- **Assumption MLR.3 (No perfect collinearity)**

„In the sample (and therefore in the population), none of the independent variables is constant and there are no exact relationships among the independent variables“

- **Remarks on MLR.3**

- The assumption only rules out perfect collinearity/correlation between explanatory variables; imperfect correlation is allowed
- If an explanatory variable is a perfect linear combination of other explanatory variables it is superfluous and may be eliminated
- Constant variables are also ruled out (collinear with intercept)

# Assumptions on MLR (3)

- Standard assumptions for the multiple regression model (cont.)
- Assumption MLR.4 (Zero conditional mean)

$E(u_i | x_{i1}, x_{i2}, \dots, x_{ik}) = 0$  ← The value of the explanatory variables must contain no information about the mean of the unobserved factors

- In a multiple regression model, the zero conditional mean assumption is much more likely to hold because fewer things end up in the error

# Assumptions on MLR (4)

- Discussion of the zero mean conditional assumption
  - Explanatory variables that are correlated with the error term are called endogenous; endogeneity is a violation of assumption MLR.4
  - Explanatory variables that are uncorrelated with the error term are called exogenous; MLR.4 holds if all explanat. var. are exogenous
  - Exogeneity is the key assumption for a causal interpretation of the regression, and for unbiasedness of the OLS estimators

- Theorem 3.1 (Unbiasedness of OLS)

$$MLR.1-MLR.4 \quad \Rightarrow \quad E(\hat{\beta}_j) = \beta_j, \quad j = 0, 1, \dots, k$$

- Unbiasedness is an average property in repeated samples; in a given sample, the estimates may still be far away from the true values

# MLR Unbiasedness

- ◆ Population model is linear in parameters:  
$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + u$$
- ◆ We can use a random sample of size  $n$ ,  
 $\{(x_{i1}, x_{i2}, \dots, x_{ik}, y_i): i=1, 2, \dots, n\}$ , from the population model, so that the sample model is  $y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik} + u_i$
- ◆  $E(u/x_1, x_2, \dots, x_k) = 0$ , implying that all of the explanatory variables are exogenous
- ◆ None of the  $x$ 's is constant, and there are no exact linear relationships among them

# Simple vs Multiple Reg Estimate

Compare the simple regression  $\tilde{y} = \tilde{\beta}_0 + \tilde{\beta}_1 x_1$   
with the multiple regression  $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2$

Generally,  $\tilde{\beta}_1 \neq \hat{\beta}_1$  unless :


$\hat{\beta}_2 = 0$  (i.e. no partial effect of  $x_2$ ) OR

$x_1$  and  $x_2$  are uncorrelated in the sample

# Including /Omitting Irrelevant Variables

- **Including irrelevant variables in a regression model**


$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + u$$

No problem because  $E(\hat{\beta}_3) = \beta_3 = 0$ .  = 0 in the population

However, including irrelevant variables may increase sampling variance.

- **Omitting relevant variables: the simple case**

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + u$$

 True model (contains  $x_1$  and  $x_2$ )

$$y = \alpha_0 + \alpha_1 x_1 + w$$

 Estimated model ( $x_2$  is omitted)

# Too Many or Too Few Variables

- What happens if we include variables in our specification that don't belong?
- There is no effect on our parameter estimate, and OLS remains unbiased
- What if we exclude a variable from our specification that does belong?
- OLS will usually be biased



# Summary of Direction Bias

	$\text{Corr}(x_1, x_2) > 0$	$\text{Corr}(x_1, x_2) < 0$
$\beta_2 > 0$	Positive bias	Negative bias
$\beta_2 < 0$	Negative bias	Positive bias

# Omitted Variable Bias Summary

- Two cases where bias is equal to zero
  - $\beta_2 = 0$ , that is  $x_2$  doesn't really belong in model
    - $x_1$  and  $x_2$  are uncorrelated in the sample
- If correlation between  $x_2$ ,  $x_1$  and  $x_2$ ,  $y$  is the same direction, bias will be positive
- If correlation between  $x_2$ ,  $x_1$  and  $x_2$ ,  $y$  is the opposite direction, bias will be negative

# The More General Case

- Technically, can only sign the bias for the more general case if all of the included  $x$ 's are uncorrelated
- Typically, then, we work through the bias assuming the  $x$ 's are uncorrelated, as a useful guide even if this assumption is not strictly true

# Goodness-of-fit: Adjusted R-square (1)

- More on goodness-of-fit and selection of regressors
- **General remarks on R-squared**
  - A high R-squared does not imply that there is a causal interpretation
  - A low R-squared does not preclude precise estimation of partial effects
- **Adjusted R-squared**
  - What is the ordinary R-squared supposed to measure?

$$R^2 = 1 - \frac{SSR}{SST} = 1 - \frac{(SSR/n)}{(SST/n)}$$



# Goodness-of-fit: Adjusted R-square (2)

- Adjusted R-squared (cont.)

- A better estimate taking into account degrees of freedom would be

$$\bar{R}^2 = 1 - \frac{(SSR/(n - k - 1))}{(SST/(n - 1))} = \text{adjusted } R^2$$

Correct degrees of freedom of nominator and denominator

- The adjusted R-squared imposes a penalty for adding new regressors
- The adjusted R-squared increases if, and only if, the t-statistic of a newly added regressor is greater than one in absolute value

- Relationship between R-squared and adjusted R-squared

$$\bar{R}^2 = 1 - (1 - R^2)(n - 1)/(n - k - 1)$$

The adjusted R-squared may even get negative



# Goodness-of-fit: Adjusted R-square (3)

- Using adjusted R-squared to choose between nonnested models
  - Models are nonnested if neither model is a special case of the other

$$rdintens = \beta_0 + \beta_1 \log(sales) + u \quad \leftarrow R^2 = .061, \bar{R}^2 = .030$$

$$rdintens = \beta_0 + \beta_1 sales + \beta_2 sales^2 + u \quad \leftarrow R^2 = .148, \bar{R}^2 = .090$$

- A comparison between the R-squared of both models would be unfair to the first model because the first model contains fewer parameters
- In the given example, even after adjusting for the difference in degrees of freedom, the quadratic model is preferred

# Incorporating Non-linearities in SLR

- Incorporating nonlinearities: Semi-logarithmic form
- Regression of log wages on years of education

$$\log(wage) = \beta_0 + \beta_1 educ + u$$

Natural logarithm of wage

- This changes the interpretation of the regression coefficient:

$$\beta_1 = \frac{\partial \log(wage)}{\partial educ} = \frac{1}{wage} \cdot \frac{\partial wage}{\partial educ} = \frac{\frac{\partial wage}{wage}}{\partial educ}$$

Percentage change of wage

... if years of education are increased by one year

# Incorporating Non-linearities in SLR

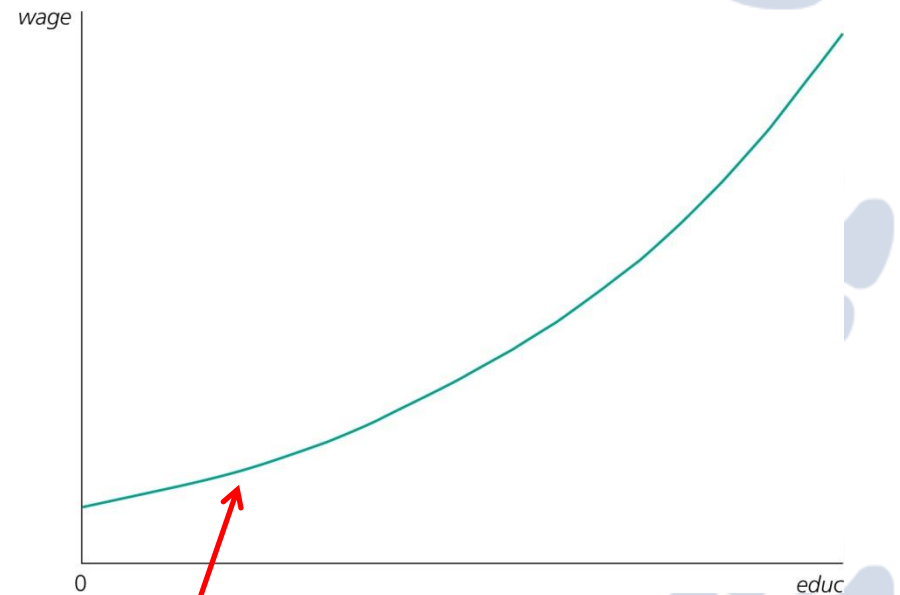
- Fitted regression

$$\widehat{\log}(wage) = 0.584 + 0.083 \text{ educ}$$

The wage increases by 8.3 % for every additional year of education (= return to education)

For example:

$$\frac{\frac{\partial wage}{wage}}{\partial educ} = \frac{\frac{+0.83\$}{10\$}}{+1 \text{ year}} = 0.083 = +8.3\%$$



Growth rate of wage is 8.3 % per year of education



# Incorporating Non-linearities in SLR

- Incorporating nonlinearities: Log-logarithmic form
- CEO salary and firm sales

$$\log(\text{salary}) = \beta_0 + \beta_1 \log(\text{sales}) + u$$

Natural logarithm of CEO salary

Natural logarithm of his/her firm's sales

- This changes the interpretation of the regression coefficient:

$$\beta_1 = \frac{\partial \log(\text{salary})}{\partial \log(\text{sales})} = \frac{\frac{\partial \text{salary}}{\text{salary}}}{\frac{\partial \text{sales}}{\text{sales}}}$$

Percentage change of salary  
... if sales increase by 1 %

Logarithmic changes are  
always percentage changes

# Introducing Quadratic Forms



- $y = \beta_0 + \beta_1 x + \beta_2 x^2$
- When  $\beta_1 > 0$ ; and  $\beta_2 < 0$ , then
- $x^* = \frac{\beta_1}{-2\beta_2}$

FIGURE A.3 Graph of  $y = 6 + 8x - 2x^2$ .

